

Original article

Temporal trends and prognostic impact of length of hospital stay in uncomplicated ST-segment elevation myocardial infarction in Spain

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ABSTRACT

Introduction and objectives: There are few data on the safety of length of stay in uncomplicated ST-segment elevation myocardial infarction. We studied trends in hospital stay and the safety of short (≤ 3 days) vs long hospital stay in Spain.**Methods:** Using data from the Minimum Basic Data Set, we identified patients with uncomplicated ST-segment elevation myocardial infarction who underwent primary percutaneous coronary intervention and who were discharged alive between 2003 and 2015. The mean length of stay was adjusted by multilevel Poisson regression with mixed effects. The effect of short length of stay on 30-day readmission for cardiac diseases was evaluated in episodes from 2012 to 2014 by propensity score matching and multilevel logistic regression. We also compared risk-standardized readmissions for cardiac diseases and mortality rates.**Results:** The adjusted length of stay decreased significantly (incidence rate ratio < 1 ; $P < .001$) for each year after 2003. Short length of stay was not an independent predictor of 30-day readmission (OR, 1.10; 95%CI, 0.92-1.32) or mortality (OR, 1.94; 95%CI, 0.93-14.03). After propensity score matching, no significant differences were observed between short and long hospital stay (OR, 1.26; 95%CI, 0.98-1.62; and OR, 1.50; 95%CI, 0.48-5.13), respectively. These results were confirmed by comparisons between risk-standardized readmissions for cardiac disease and mortality rates, except for the 30-day mortality rate, which was significantly higher, although probably without clinical significance, in short hospital stays (0.103% vs 0.109%; $P < .001$).**Conclusions:** In Spain, hospital stay ≤ 3 days significantly increased from 2003 to 2015 and seems a safe option in patients with uncomplicated ST-segment elevation myocardial infarction.

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Tendencias e impacto pronóstico de la duración de la estancia hospitalaria en el infarto de miocardio con elevación del segmento ST no complicado en España

RESUMEN

Introducción y objetivos: La información sobre la seguridad de la duración de la estancia es escasa en el infarto de miocardio con elevación del segmento ST no complicado. Se han estudiado las tendencias y la seguridad en España de la estancia corta (≤ 3 días) frente a la prolongada.**Métodos:** Se identificaron en el Conjunto Mínimo Básico de Datos los episodios de pacientes con infarto de miocardio con elevación del segmento ST no complicado tratados con intervención coronaria percutánea primaria y dados de alta vivos entre 2003 y 2015. La estancia media se ajustó mediante regresión de Poisson multinivel con efectos mixtos. El efecto de la estancia corta en el reingreso por causa cardiovascular a 30 días se evaluó en episodios de 2012-2014 mediante emparejamiento por puntuaciones de propensión y regresión logística multinivel, comparando las razones estandarizadas de reingreso y mortalidad por riesgo.**Resultados:** La estancia ajustada disminuyó significativamente (razón de tasas de incidencia < 1 ; $p < 0,001$) cada año desde 2003. La estancia corta no fue un predictor independiente de reingreso (OR = 1,10; IC95%, 0,92-1,32) ni de mortalidad (OR = 1,94; IC95%, 0,93-14,03). Después del emparejamiento, tampoco hubo diferencias significativas en ambos casos (OR = 1,26; IC95%, 0,98-1,62; y OR = 1,50; IC95%, 0,48-5,13). Las comparaciones entre las razones estandarizadas de reingreso y mortalidad por riesgo confirmaron estos resultados, excepto en la de mortalidad a los 30 días,

Palabras clave:

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significativamente mayor en la estancia corta, aunque probablemente sin significado clínico (el 0,103 y el 0,109%; $p < 0,001$).

Conclusiones: La estancia ≤ 3 días aumentó significativamente en España desde 2003 a 2015 y parece una opción segura en el infarto de miocardio con elevación del segmento ST no complicado.

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Abbreviations

MBDS: Minimum Basic Data Set
 pPCI: primary percutaneous coronary intervention
 SNS: Spanish National Health System
 STEMI: ST-segment elevation myocardial infarction

INTRODUCTION

Primary percutaneous coronary intervention (pPCI) is the ideal reperfusion strategy for ST-segment elevation myocardial infarction (STEMI).¹ The pPCI rate has been boosted through the implementation of health care networks, reducing complications and shortening hospital stay.^{1–4} The current clinical practice guidelines of the European Society of Cardiology consider a short length of hospital stay (48–72 hours) to be a valid option for uncomplicated STEMI.¹ However, this recommendation is mainly based on older studies with small sample sizes,^{3,5–10} which undermines reliable analyses of the safety effects of a short length of stay.

In Spain, the recent length of stay trends in STEMI patients treated with pPCI are unknown and no information is available on differences among hospitals and there are no data supporting the recommendation for a short length of stay.¹

Thus, the aim of the present study was to evaluate the temporal trends in length of stay in patients admitted to Spanish National Health System (SNS) hospitals with uncomplicated STEMI treated with PCI and its possible association with adjusted rates of 30-day risk of death and readmission for cardiovascular diseases in these patients. The results will help to determine whether the recommendation of a short length of stay for these uncomplicated events is safe in the Spanish SNS and identify possible improvements to the health care process for STEMI in Spain.

METHODS

Study design, data source, and patient population

This retrospective observational study investigated patients admitted for STEMI to SNS hospitals and treated with PCI. The data source was the Minimum Basic Data Set (MBDS)¹¹ of the Ministry of Health, Consumer Affairs, and Social Welfare. The usefulness of this database for the study of acute coronary syndrome in Spain was recently validated.¹² The following events occurring between January 1, 2003, and December 31, 2005, were selected: patients with a principal diagnosis of STEMI who were treated with PCI and discharged alive (ICD-9-CM codes; [table 1 of the supplementary data](#)). Patients treated with thrombolysis or cardiac surgery were excluded to reduce analytical biases. Also excluded were patients with any of the following complications associated with the index event: heart failure, acute pulmonary edema, ventricular arrhythmias, cardiac arrest, and cardiogenic shock ([table 1 of the supplementary data](#)).

To improve the consistency and quality of the data, a) events corresponding to patients admitted to one hospital and transferred to another for PCI were concatenated and the lengths of the stay in

the referring hospital were attributed to the second hospital, and b) events corresponding to patients younger than 35 years and older than 94 years were excluded, as well as discharges against medical advice or due to transfer to a social health center or an unknown reason, patients with events not causing hospitalization who were discharged to home, patients transferred to another hospital who returned to the referring hospital, and events classified in Major Diagnostic Category 14 (pregnancy, childbirth, and the puerperium) of the All Patient Refined Diagnosis-Related Groups.¹³

The length of hospital stay, measured in days, was calculated as the difference between the discharge and admission dates and its changes over time were analyzed between 2003 and 2015. The impact of a short length of hospital stay (≤ 3 days) was evaluated for the 2012 to 2014 period because the MBDS enabled reliable identification of readmission events during this period. Outcome variables were unscheduled readmission 30 days after discharge from the index event for cardiovascular diseases (rheumatic heart disease, hypertensive heart disease, ischemic heart disease, pulmonary vascular diseases, other heart diseases, and aortic dissection and aneurysm, as well as other admissions to cardiology and cardiac surgery, regardless of the principal diagnosis) and the in-hospital mortality during these readmissions (because the MBDS does not include out-of-hospital information).

Statistical analysis

Multilevel mixed-effects Poisson regression was applied to adjust the length of stay, given its skewed distribution.¹⁴ In addition to year of discharge and patient age and sex, the risk factors included in the models designed by the Centers for Medicare and Medicaid Services (CMS methodology) were used to adjust for the risk of mortality and readmission for AMI.^{15,16} We adapted the models to the structure of the MBDS, after grouping the secondary diagnoses according to the clinical condition categories proposed by Pope et al.,¹⁷ updated annually by the Agency for Healthcare Research and Quality.¹⁸ The expected length of stay was derived from the individual predictions obtained from the estimated model.

Taking into account the existence of characteristics specific to the patients and to the treating centers that are independent of the quality of the health care provided,¹⁹ the 30-day readmission and mortality rates were adjusted for risk according to the CMS methodology. The independent variables were those included in the models for readmission and mortality from AMI, respectively, and a dichotomous variable was included to identify whether the stay was short or long.

Multilevel logistic regression models²⁰ were constructed that, in addition to clinical and demographic variables, included a specific random effect related to the hospital^{21,22}; using backward elimination, significance levels of $P < .05$ and $P \geq .10$ were applied for factor selection and elimination, respectively. The discrimination of the definitive models was assessed using the area under the receiver operating characteristic curve (AUC).

In all multilevel models, if the treating hospital could not be identified, the corresponding events were eliminated; in addition, the incidence rate ratios or odds ratios (ORs) were calculated, as appropriate, as well as their 95% confidence intervals (95% CIs).

The risk-standardized readmission and mortality ratios (RSRRs and RSMRs, respectively) were calculated as the ratios between the expected outcome (which individually considers the functioning of the hospital treating the patient) and the expected outcome (which considers standard functioning according to the average of all hospitals) multiplied by the crude readmission or mortality rate of the study population.^{18,23} If the RSRR (or RSMR) of a hospital is greater than the crude readmission rate (or crude mortality rate), the probability of readmission (or mortality) in that hospital is higher than the average of the hospitals studied.

To minimize selection bias, the impact of a short length of stay on readmission and mortality during these readmissions was assessed with propensity score matching (k-nearest neighbors matching option, psmatch2, Stata). To do this, we selected long-stay events that, based on the patients' demographic characteristics and comorbidities, had a similar short-stay probability to that of each short stay event. The matching was performed using the risk adjustment models at a 1:1 ratio and with a maximum propensity score difference (caliper) of 0.05 standard deviations and without replacement. Also calculated were the probability of readmission or death, the effect of the differences between the 2 groups (average treatment effect on the treated [ATT]), and the ORs with their 95% CIs.

To discriminate between high- and low-volume hospitals (according to the number of events attended), a k-means clustering algorithm was used in two thirds of the dataset to obtain the maximum intracluster and minimum intercluster densities, which were validated with the remaining third of the dataset.

The RSRR and RSMR of 30-day readmissions were compared between the 2 length of stay groups classified according to their complexity using the RECALCAR²⁴ typology (table 2 of the supplementary data) and the volume of events recorded.

As sensitivity analysis, the impact of transfers between hospitals was evaluated by excluding index events with discharge due to transfer.

Quantitative variables are presented as means ± standard deviations and categorical variables as frequencies and percentages. Correlations among quantitative variables were analyzed with the Spearman rank coefficient (ρ). The t test was used to compare 2 categories, whereas analysis of variance (ANOVA) with the Bonferroni correction was used for 3 or more. Comparisons among discrete variables were made using the chi-square test or Fisher exact test. All comparisons were 2-sided and differences were considered significant at *P* < .05. All statistical analyses were performed with STATA 13 or SPSS 21.0.

RESULTS

Temporal trends in length of hospital stay

In total, 205 016 hospitalization events with STEMI as the principal diagnosis and treatment with PCI were identified; of these, 190 078 patients were discharged alive. Once transfer events were concatenated across hospitals, 188 854 events remained. After exclusions, 134 002 events comprised the study population for the analysis of trends (figure 1).

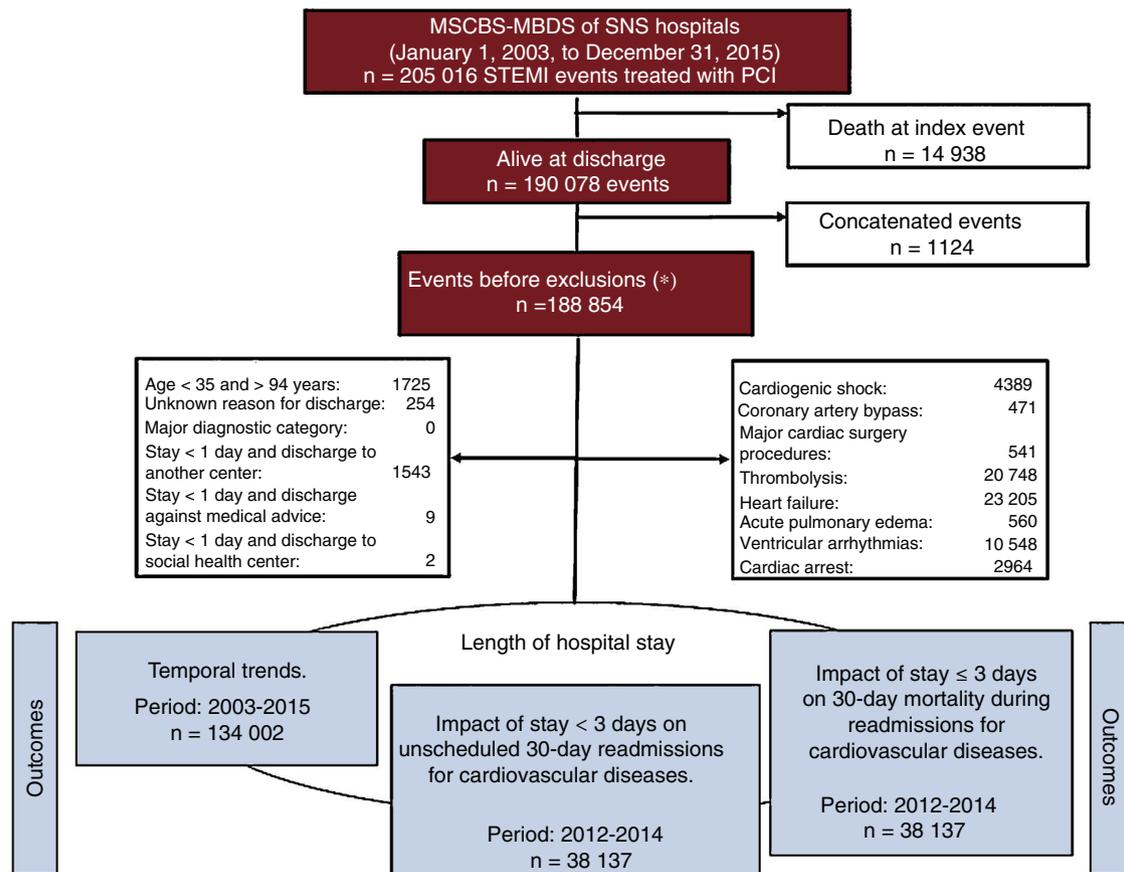


Figure 1. Flowchart of patients' events. MSCBS-MBDS, Minimum Basic Data Set provided by the Spanish Ministry of Health, Consumer Affairs, and Social Welfare; PCI, percutaneous coronary intervention; SNS, Spanish National Health System; STEMI, ST-segment elevation myocardial infarction. *Exclusions were not mutually exclusive.

The crude mean length of hospital stay was 7.2 ± 5.5 days and ranged between an annual maximum of 9.0 ± 7.2 days in 2003 and a minimum of 6.1 ± 4.7 days in 2015. The length decreased each year vs the previous year by 3.2%. The adjusted length of hospital stay was significantly reduced (incidence rate ratio < 1 ; $P < .001$) each year of the period analyzed. With 2003 as reference (table 1), the expected mean lengths of stay showed significant annual differences (figure 2) and the percentage of short lengths of stays increased (from 14.3% in 2003 to 19.3% in 2015; $P < .001$) (figure 3).

Impact of a short length of stay on readmissions and mortality

Between January 1, 2012, and December 31, 2014, 38 137 index events (28.5%) were identified from all patients with STEMI index events treated with PCI recorded in the MBDS. Of these, 6486 (17%) were short. Short lengths of stay were more likely in women, younger people, and patients with fewer comorbidities (table 2).

The most frequent causes of readmission are shown in table 3 of the supplementary data. The crude rate of 30-day readmission was

Table 1
Adjusted model for the length of hospital stay

	IRR	95%CI	P
<i>Year of study</i>			
2003 (reference)	1.00	—	—
2004	0.94	0.92-0.97	< .001
2005	0.90	0.86-0.94	< .001
2006	0.86	0.82-0.90	< .001
2007	0.83	0.79-0.88	< .001
2008	0.80	0.75-0.85	< .001
2009	0.77	0.72-0.82	< .001
2010	0.74	0.69-0.78	< .001
2011	0.70	0.66-0.74	< .001
2012	0.68	0.63-0.72	< .001
2013	0.66	0.62-0.71	< .001
2014	0.64	0.60-0.68	< .001
2015	0.62	0.58-0.66	< .001
<i>Female sex</i>	1.04	1.03-1.05	< .001
<i>Age (per 1-y increment)</i>	1.004	1.003-1.004	< .001
<i>DM or DM complications</i>	1.05	1.04-1.06	< .001
<i>Stroke</i>	1.87	1.65-2.11	< .001
<i>Cerebrovascular disease</i>	1.10	1.05-1.15	< .001
<i>History of vascular disease/complications</i>	2.05	1.63-2.58	< .001
<i>Hemiplegia, paraplegia, paralysis, functional disability</i>	1.16	1.10-1.22	< .001
<i>Vascular or circulatory disease</i>	1.13	1.10-1.16	< .001
<i>History of acute myocardial infarction</i>	1.17	1.07-1.28	.001
<i>History of other acute/subacute ischemic heart diseases</i>	1.13	1.10-1.15	< .001
<i>History of congestive heart failure</i>	1.10	1.06-1.13	< .001
<i>History of cardiorespiratory failure or shock (noncardiogenic)</i>	1.65	1.52-1.79	< .001
<i>History of specific arrhythmias and other heart rhythm disorders</i>	1.14	1.12-1.15	< .001
<i>Valvular or rheumatic heart disease</i>	1.08	1.06-1.11	< .001
<i>Chronic obstructive pulmonary disease</i>	1.03	1.01-1.05	< .001
<i>Asthma</i>	1.05	1.01-1.09	.024
<i>History of pneumonia</i>	1.37	1.31-1.44	< .001
<i>History of cancer</i>	1.13	1.09-1.18	< .001
<i>History of metastatic cancer or acute leukemia</i>	0.88	0.77-1.00	.042
<i>Major trauma in the previous year</i>	1.47	1.31-1.65	< .001
<i>Severe psychiatric disorders</i>	1.11	1.06-1.17	< .001
<i>Dementia and other specific brain disorders</i>	1.28	1.20-1.37	< .001
<i>Anterior myocardial infarction</i>	0.89	0.83-0.96	.003
<i>Myocardial infarction in other locations</i>	0.89	0.83-0.95	.001
<i>Protein-energy malnutrition</i>	1.69	1.13-2.54	.011
<i>Electrolyte or acid-base disorders</i>	1.17	1.12-1.22	< .001
<i>Iron deficiency or other anemias and specific blood diseases</i>	1.28	1.24-1.31	< .001
<i>Chronic kidney disease</i>	1.25	1.22-1.28	< .001
<i>Other urinary tract disorders</i>	1.25	1.19-1.31	< .001
<i>Chronic decubitus or cutaneous ulcer</i>	1.95	1.58-2.41	< .001

95%CI, 95% confidence interval; DM, diabetes mellitus; IRR, incidence rate ratio.
N = 132 715; 1287 (0.96%) lost events.

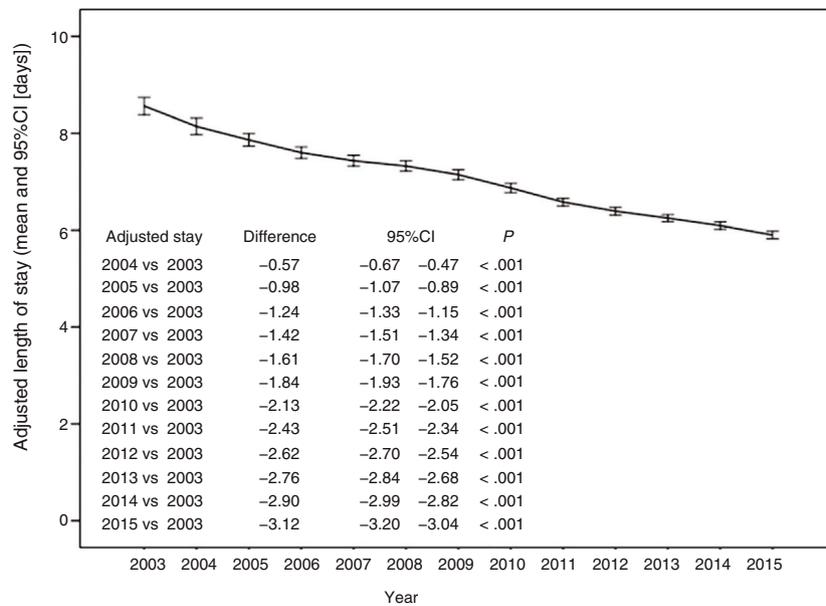


Figure 2. Annual changes in mean expected lengths of hospital stay (adjusted for the length the stay using Poisson multilevel regression). The bars represent the 95%CI of the mean adjusted length of stay. 95%CI, 95% confidence interval.

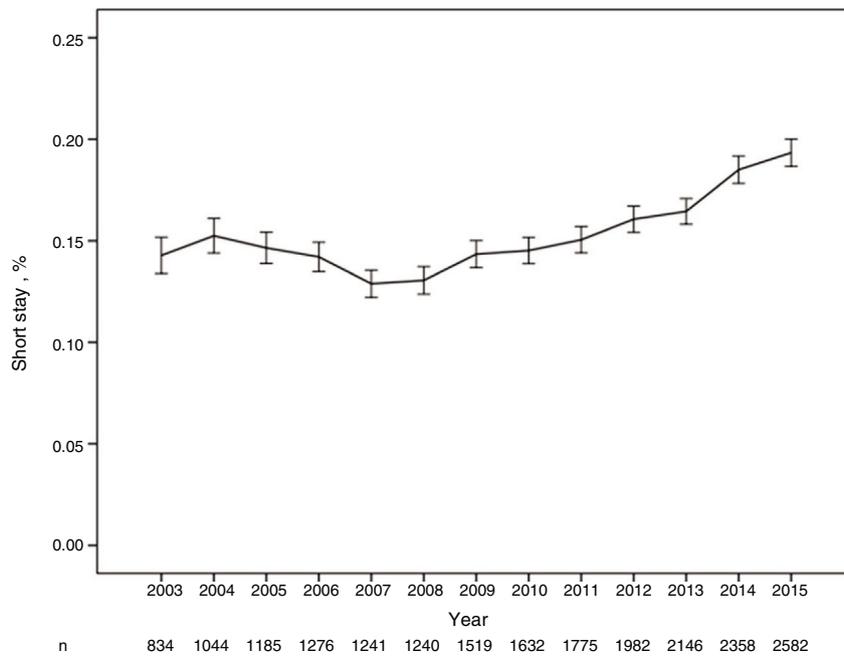


Figure 3. Annual changes in the proportions of discharges from short (≤ 3 days) hospital stays.

2.59% (2.69% for short hospital stays vs 2.61%) and the crude rate of mortality during these readmissions was 0.11% (0.15% vs 0.10%).

The adjusted risk models of 30-day readmissions and mortality during these admissions are shown in table 3. A short hospital stay was not an independent predictor in either of the models: OR = 1.10 (95%CI, 0.92-1.32; $P = .30$) and OR = 1.94 (95%CI, 0.93-14.03; $P = .077$), respectively.

Discrimination for 30-day readmission was low (AUC = 0.65) but was appreciably higher (AUC = 0.84) for mortality during these readmissions (figure 1 of the supplementary data). The intraclass correlation coefficients of both models were 0.02 and 0.11,

respectively, and their median ORs were 1.3 and 1.8, indicating considerable variability among hospitals.

To analyze the effects of a short hospital stay, propensity score matching was used to obtain 2 cohorts of 12 966 events for readmissions and of 12 972 events for mortality (99.99% and 100% of the events with short hospital stays). The characteristics of these groups are shown in table 4. After matching, there were no significant differences between the 2 groups in 30-day readmissions (ATT = 0.024 vs 0.019; $P = .204$; OR = 1.26; 95%CI, 0.98-1.62) or 30-day mortality (ATT = 0.001 vs 0.001; $P = .52$; OR = 1.5; 95%CI, 0.48-5.13).

Table 2
Profile of the patients with index events recorded from 2012 to 2014

	Short stay (≤ 3 d) n=6486	Long stay (> 3 d) n=31 651	P
Age, y	60.6 \pm 12.2	63.1 \pm 12.7	< .001
Hospital stay, d	2.4 \pm 0.6	7.3 \pm 4.6	< .001
Female sex, %	18.2	21.4	< .001
Hypertension and hypertensive heart disease, %	46.7	51.2	< .001
DM or DM complications, %	21.4	25.3	< .001
History of arrhythmias, %	12.0	17.2	< .001
Valvular or rheumatic heart disease, %	5.1	9.2	< .001
Vascular or circulatory disease, %	5.2	7.8	< .001
Chronic obstructive pulmonary disease, %	5.7	6.5	< .001
Chronic kidney disease, %	3.3	6.6	< .001
Congestive heart failure, %	4.4	6.0	< .001
Iron deficiency or other anemias and specific blood diseases, %	1.7	4.0	< .001
Anterior myocardial infarction, %	2.6	3.6	< .001

DM, diabetes mellitus.

Total number of patients = 38 137.

Table 3
Independent predictors of 30-day readmission after discharge and mortality during the readmission

	30-d readmissions		Mortality during the 30-d readmission	
	OR (95%CI)	P	OR (95%CI)	P
Female sex	1.18 (1.02-1.38)	.03	—	—
Age (per 1-y increment)	1.02 (1.01-1.02)	< .001	1.06 (1.03-1.01)	< .001
Short stay (≤ 3 d)	1.10 (0.92-1.32)	.30	1.94 (0.93-14.03)	.077
History of congestive heart failure	1.80 (1.45-2.23)	< .001	—	—
Valvular or rheumatic heart disease	1.43 (1.18-1.74)	< .001	2.99 (1.50-6.00)	.002
History of specific arrhythmias and other heart rhythm disorders	1.29 (1.10-1.51)	.002	—	—
Chronic kidney disease	1.58 (1.28-1.95)	< .001	—	—
Hemiplegia, paraplegia, paralysis, functional disability	—	—	2.99 (1.50-6.00)	.018
Chronic liver disease	—	—	4.27 (1.30-14.05)	.03
Constant	0.005 (0.004-0.008)	< .001	0.001 (0.00-0.001)	< .001

95%CI, 95% confidence interval; OR, odds ratio.

There were also no significant differences between the 2 groups in the RSRR (2.54% vs 2.55%; $P = .020$). Although the mean RSMR at 30 days was significantly higher in the short hospital stay group, the difference was probably not clinically significant (0.103% vs 0.109%; $P < .001$) (table 4 of the supplementary data).

Impact of patient volume and type of hospital

From 2012 to 2014, hospitals with higher volume (> 415 events) had a larger proportion of short hospital stays (18.10% vs 17.10%; $P = .02$), but no significant differences were found in the mean lengths of stay, the 30-day RSRR, or the 30-day RSMR according to volume (all $P > .05$). Neither were there significant correlations of the RSMR or RSRR with the patient volume nor were their differences significant according to hospital type (all $P > .05$).

Sensitivity analysis

In the sensitivity analysis, events corresponding to 4192 patients (11%) from 2012 to 2014 were excluded due to their transfer to another hospital. As in the original models, neither a short hospital stay nor their respective ORs or ATTs obtained from matching were significant.

DISCUSSION

The main finding of this study is that, with a very large population (that of the Spanish SNS) and an extended period of time (2003-2015), hospitalization of patients with uncomplicated STEMI for ≤ 3 days is practically safe. Our results substantiate the decision of the European clinical practice guidelines to upgrade the level of recommendation for this practice to IIa.¹ Another notable finding is the progressive and significant decrease in the mean length of hospital stay in patients admitted for STEMI during the study period, which probably reflects clinical strategy improvements.

After matching, which enables a less biased analysis than RSRR and RSMR comparisons due to an appropriate balance between the short and long hospital stay groups, there were no significant differences.

The RSRR for cardiovascular diseases was also similar in the 2 groups and, although the 30-day RSMR in the short stay group was higher (0.109% vs 0.103%), the difference does not seem clinically relevant because it is equivalent to just 6 deaths/100 000 patients. This represents an excess mortality rate < 1 death per year associated with a short hospital stay.

Studies largely performed in the United States have also found a significant decrease in length of hospital stay in recent years.^{3,4,25} However, in the Spanish SNS, the tendencies in the length of

Table 4

Clinical and demographic characteristics of the patients in the short and long hospital stay groups before and after propensity score matching

	Before matching			After matching		
	Short stay (≤ 3 d)	Long stay (> 3 d)	P	Short stay (≤ 3 d)	Long stay (> 3 d)	P
<i>Unscheduled 30-d readmissions for cardiovascular disease</i>						
Number of valid cases	6486	31 615		6483	6483	
Mean age, y	60.6	63.0	< .001	60.6	60.7	.796
Female sex, %	18.2	21.7	< .001	18.2	18.4	.89
Electrolyte imbalance, %	0.9	1.6	< .001	0.9	0.8	.45
History of congestive heart failure, %	4.4	6.0	< .001	4.4	4.2	.70
Acute coronary syndrome, %	4.9	6.9	< .001	4.9	4.9	.90
Valvular or rheumatic heart disease, %	6.1	9.2	< .001	6.1	5.0	.84
History of specific arrhythmias and other heart rhythm disorders, %	12.0	17.2	< .001	12.0	12.1	.87
Renal failure, %	3.3	6.6	< .001	3.3	3.3	.84
<i>Mortality during the 30-d readmission</i>						
Number of valid cases	31 651	6486		6486	6486	
Mean age, y	60.6	63.0	< .001	60.6	60.7	.86
History of coronary revascularization surgery, %	0.8	0.8	.81	0.8	0.9	.92
Valvular or rheumatic heart disease, %	5.1	9.0	< .001	5.1	5.1	.94
Hemiplegia, paraplegia, paralysis, functional disability	0.9	1.3	.006	0.9	0.8	.64
Chronic liver disease, %	0.3	0.3	.56	0.3	0.3	.37

n, number of valid cases.

hospital stay for STEMI and its prognostic impact had not been characterized, and the availability of data on the impact of a short hospital stay could help to improve the health care process.

In the CathPCI registry,²⁵ which included 33 920 patients with STEMI treated with pPCI, the proportion of hospital stays ≤ 3 days increased from 24% in 2004 to 30% in 2009, a percentage increase similar to that found in Spain in the 2003 to 2015 period (5%; from 14.30% in 2003 to 19.30% in 2015), although the percentage of stays ≤ 3 days was considerably lower. In a study with more than 50 000 patients, a significantly longer hospital stay was seen (at least 3–4 days) in European countries, including Spain.²⁶ Potentially unnecessary days consumed per 100 patients enrolled ranged between 65 (New Zealand) and 839 (Germany) and the potential for a more efficient hospital stay for low-risk patients was suggested to be particularly marked in European countries. Independently of the causes of these differences, the safety of a short hospital stay for STEMI should underline the significance of an efficient hospital stay. In this regard, the information on the efficiency and safety of a shortened hospital stay for STEMI patients is mainly derived from observational studies,^{3,6,9,10,25} and few were recent and multicenter.^{9,25} A recent meta-analysis concluded that a hospital stay ≤ 3 days does not increase the risk of 30-day or 6-month readmission or death.²⁷ However, because the analysis included studies from 1998 to 2016 and had a small sample size, heterogeneous definition of low-risk STEMI, and varying follow-up protocols after discharge, its conclusions are limited.

The reduced length of stay between 2003 and 2015 in our study could be due to more widespread use of PCI and to other therapeutic and STEMI treatment-related improvements and advances, as well as other factors, such as the more efficient clinical management of these patients. However, because the proportion of patients with a hospital stay > 3 days was very high (> 80%), a short hospital stay can be considered uncommon for low-risk patients in Spain. A short hospital stay is a valid alternative for low-risk patients with STEMI,^{9,10,25} who may represent up to 50% to 70% of all patients with STEMI.^{5,6,9,25,26} De Luca et al.⁶ observed that a longer length of stay in low-risk STEMI would only save, at 30 days after discharge, 1 life per 1097 patients, at an additional cost of almost €200 000. In

a cost-effectiveness study, Newby et al.²⁸ reported that only 0.006 years of life per patient was saved when patients were hospitalized for an additional (fourth) day. In this regard, we believe that an early monitoring program, particularly in cardiac rehabilitation units, could help to optimize the length of hospital stay. Regardless, the objective of this study was not to explore the estimated saving that a shortened hospital stay could represent, but to reflect the reality of the health care administered in terms of hospital stay and its effect on prognosis. A shorter hospital stay in low-risk patients can reduce complications and costs.²⁹ However, a potential disadvantage could be related to modification of the patients' perception of the severity of the disease, which could undermine efforts to improve secondary prevention.¹

Limitations

Although this retrospective analysis is based on administrative data, the validity of administrative registries is comparable to that of medical registries^{12,22} and the reliability of these studies allows public comparison of hospitals in terms of outcomes.³⁰

In addition, in contrast to the CMS methodology, only readmissions due to cardiovascular disease were analyzed (and mortality during these readmission) due to a lack of events with a principal diagnosis other than that of cardiovascular disease (not included in the MBDS provided by the Spanish Ministry of Health, Consumer Affairs, and Social Welfare). However, cardiac events are probably a better indicator of performance related to length of stay than all-cause readmissions.³¹ Moreover, the difference in the mortality of patients with STEMI from cardiovascular diseases observed in other studies^{32,33} vs our findings is likely due to the absence of patient selection in those studies. For example, from our population with STEMI, we excluded patients with heart failure, acute pulmonary edema, ventricular arrhythmias, cardiac arrest, and cardiogenic shock. In addition, readmissions were not considered a recurrent outcome variable. However, the risk of adverse events was low or very low in our population, and a low burden of recurrences would be expected.

Another limitation lies in the inability of ICD-9-CM coding to accurately determine whether primary PCI was performed. However, according to the Spanish Cardiac Catheterization and Coronary Intervention Registry,³⁴ 86% of PCIs in STEMI are primary procedures and 5% are rescue PCIs (after fibrinolysis, which were excluded from our study population). Thus, 91% of the PCIs would be primary in our study.

CONCLUSIONS

Hospital stay for uncomplicated STEMI treated with PCI significantly decreased in Spain between 2003 and 2015, although most patients are still hospitalized for 4 or more days. The discharge of these patients in ≤ 3 days can be considered safe in the SNS and its generalization, in line with the recommendations of the European clinical practice guidelines, would improve the efficient use of health care resources.

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CONFLICTS OF INTEREST

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WHAT IS KNOWN ABOUT THE TOPIC?

- The increased rate of percutaneous coronary intervention in STEMI has helped to reduce complications and shorten length of hospital stay.
- Although the recommendation has recently been strengthened for discharge in less than 3 days in low-risk patients with STEMI, the scientific evidence is slight overall and entirely lacking in the Spanish SNS.

WHAT DOES THIS STUDY ADD?

- In Spain, hospital stay in STEMI patients significantly shortened from 2003 to 2015. The percentage of short stays (≤ 3 days) significantly increased from 14.30% (2003) to 19.30% (2015).
- After propensity score matching, there were no differences in 30-day readmission or mortality during the readmission between a short and a long stay.

APPENDIX. SUPPLEMENTARY DATA

Supplementary data associated with this article can be found in the online version at <https://doi.org/10.1016/j.rec.2019.09.016>

REFERENCES

1. Ibáñez B, James S, Agewall S, et al. Guía ESC 2017 sobre el tratamiento del infarto agudo de miocardio en pacientes con elevación del segmento ST. *Rev Esp Cardiol.* 2017;70:1082e1-1082.e61.
2. Cequier Fillat A, Ariza-Solé A, Elola FJ, et al. Impacto sobre la mortalidad de diferentes sistemas de asistencia en red en el tratamiento del infarto agudo de miocardio con elevación del segmento ST. La experiencia de España. *Rev Esp Cardiol.* 2017;70:155–161.
3. Berger AK, Duval S, Jacobs Jr DR, et al. Relation of length of hospital stay in acute myocardial infarction to postdischarge mortality. *Am J Cardiol.* 2008;101:428–434.
4. Spencer FA, Lessard D, Gore JM, et al. Declining length of hospital stay for acute myocardial infarction and postdischarge outcomes. *Arch Intern Med.* 2004;164:733–740.
5. Grines CL, Marsalese DL, Brodie B, et al. Safety and cost-effectiveness of early discharge after primary angioplasty in low risk patients with acute myocardial infarction. *J Am Coll Cardiol.* 1998;31:967–972.
6. De Luca G, Suryapranata H, Van't Hof AW, et al. Prognostic assessment of patients with acute myocardial infarction treated with primary angioplasty: implications for early discharge. *Circulation.* 2004;109:2737–2743.
7. Azzalini L, Sole E, Sans J, et al. Feasibility and safety of an early discharge strategy after low-risk acute myocardial infarction treated with primary percutaneous coronary intervention. *Cardiology.* 2015;130:120–129.
8. Melberg T, Jorgensen M, Orn S, Solli T, Edland U, Dickstein K. Safety and health status following early discharge in patients with acute myocardial infarction treated with primary PCI: a randomized trial. *Eur J Prev Cardiol.* 2015;22:1427–1434.
9. Noman A, Zaman AG, Schechter C, Balasubramaniam K, Das R. Early discharge after primary percutaneous coronary intervention for ST-elevation myocardial infarction. *Eur Heart J Acute Cardiovasc Care.* 2013;2:262–269.
10. Jones DA, Rathod KS, Howard JP, et al. Safety and feasibility of hospital discharge 2 days following primary percutaneous intervention for ST-segment elevation myocardial infarction. *Heart.* 2012;98:1722–1727.
11. Sistema Nacional de Salud. Registro de altas de hospitalización: CMBD del Sistema Nacional de Salud. Glosario de términos y definiciones. Portal estadístico del SNS [updated May 2018]. p. 5-6. Available at: <https://estadistico.inteligenciadegestion.mscbs.es/publicoSNS/comun/DescargaDocumento.aspx?IdNodo=6415>. Accessed 28 Aug 2019.
12. Bernal JL, Barrabés JA, Íñiguez A, et al. Clinical and administrative data on the research of acute coronary syndrome in Spain: minimum basic data set validity. *Rev Esp Cardiol.* 2018;72:56–62.
13. Averill RF, McCullough EC, Goldfield N, Hughes JS, Bonazelli J, Bentley L. *3M APR DRG Classification System*. 31st ed. Wallingford, CT: Agency for Healthcare Research and Quality; 2013. Report No.: GRP-041. Available at: https://www.hcup-us.ahrq.gov/db/nation/nis/grp031_aprdrgr_meth_ovrview.pdf. Accessed 28 Aug 2019.
14. Damrauer S, Gaffey AC, DeBord SM, et al. Comparison of risk factors for length of stay and readmission following lower extremity bypass surgery. *J Vasc Surg.* 2015;62:1192–1200.
15. Centers for Medicare and Medicaid Services. Measures updates and specifications: 2017 condition-specific measures updates and specifications report hospital-level 30-day risk-standardized readmission measures. Acute myocardial infarction, chronic obstructive pulmonary disease, heart failure, pneumonia and stroke. Yale New Haven Health Services Corporation, Centers for Medicare and Medicaid Services (CMS); 2017.
16. Centers for Medicare and Medicaid Services. Measures Updates and Specifications: 2017 Condition-specific measures updates and specifications report hospital-level 30-day risk-standardized mortality measures. Yale New Haven Health Services Corporation, Centers for Medicare and Medicaid Services (CMS); 2017.
17. Pope GC, Ellis RP, Ash AS, et al. Principal inpatient diagnostic cost group model for Medicare risk adjustment. *Health Care Financ Rev.* 2000;21:93–118.
18. AHRQ QITM Version v6.0 ICD9CM. Inpatient quality indicators #91, technical specifications, mortality for selected conditions. Available at: www.qualityindicators.ahrq.gov. Accessed 15 Aug 2018.
19. Iezzoni LI. Dimensions of risk. In: Iezzoni LI, ed. *Risk adjustment for measuring health care outcomes* 2.ª ed. Ann Arbor: Health Administration Press; 1997. p. 43168.
20. Krumholz HM, Wang Y, Mattera JA, et al. An administrative claims model suitable for profiling hospital performance based on 30 day mortality rates among patients with an acute myocardial infarction. *Circulation.* 2006;113:168392.
21. Normand SLT, Glickman ME, Gatsonis CA. Statistical methods for profiling providers of medical care: issues and applications. *J Am Stat Assoc.* 1997;92:80314.
22. Goldstein H, Spiegelhalter DJ. League tables and their limitations: statistical aspects of institutional performance. *J Royal Stat Soc.* 1996;159:385444.
23. Shahian DM, Normand SL, Torchiana DF, et al. Cardiac surgery report cards: comprehensive review and statistical critique. *Ann Thorac Surg.* 2001;72:215568.
24. Íñiguez Romo A, Bertomeu Martínez V, Rodríguez Padial L, et al. The RECALCAR Project. Healthcare in the cardiology units of the Spanish National Health System, 2011 to 2014. *Rev Esp Cardiol.* 2017;70:567–575.
25. Swaminathan RV, Rao SV, McCoy LA, et al. Hospital length of stay and clinical outcomes in older STEMI patients after primary PCI: a report from the National Cardiovascular Data Registry. *J Am Coll Cardiol.* 2015;65:1161–1171.
26. Kaul P, Newby LK, Fu Y, et al. International differences in evolution of early discharge after acute myocardial infarction. *Lancet.* 2004;363:511–517.
27. Gong W, Li A, Ai H, Shi H, Wang X, Nie S. Safety of early discharge after primary angioplasty in low-risk patients with ST-segment elevation myocardial infarction: A meta-analysis of randomised controlled trials. *Eur J Prev Cardiol.* 2018;25:807–815.

28. Newby LK, Hasselblad V, Armstrong PW, et al. Time-based risk assessment after myocardial infarction. Implications for timing of discharge and applications to medical decision-making. *Eur Heart J*. 2003;24:182–189.
29. Zhan C, Miller MR. Excess length of stay, charges, and mortality attributable to medical injuries during hospitalization. *JAMA*. 2003;290:1868–1874.
30. Centers for Medicare and Medicaid Services. Specifications Manual for National Hospital Inpatient Quality Measures, Version 3.1 a. Centers for Medicare and Medicaid Services (CMS), The Joint Commission. Apr 1, 2010. Available at: https://www.joint-commission.org/assets/1/18/ReleaseNotes_3.1a.pdf. Accessed 28 Aug 2019.
31. Southern DA, Ngo J, Martin BJ, et al. Characterizing types of readmission after acute coronary syndrome hospitalization: implications for quality reporting. *J Am Heart Assoc*. 2014;3:e001046.
32. Pedersen F, Butrymovich V, Kelbæk H, et al. Short- and long-term cause of death in patients treated with primary PCI for STEMI. *J Am Coll Cardiol*. 2014;64:2101–2108.
33. Yamashita Y, Shiomi H, Morimoto T, et al. Cardiac and noncardiac causes of long-term mortality in ST-segment-elevation acute myocardial infarction patients who underwent primary percutaneous coronary intervention. *Circ Cardiovasc Qual Outcomes*. 2017. <http://dx.doi.org/10.1161/CIRCOUTCOMES.116.002790>.
34. Cid Álvarez AB, Rodríguez Leor O, Moreno R, Pérez de Prado A. Spanish Cardiac Catheterization and Coronary Intervention Registry. 27th Official Report of the Spanish Society of Cardiology Working Group on Cardiac Catheterization and Interventional Cardiology (1990–2017). *Rev Esp Cardiol*. 2018;71:1036–1046.